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Interpretable variable selection and differential structure estimation in graphical models

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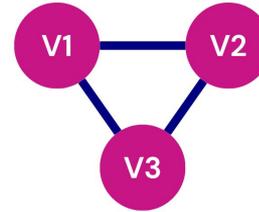
	Supporting Structure	Platforms (log ₁₀ order of magnitude)	Features
 Genome	DNA	Microarrays Sequencing	Categorical data Distance-driven correlation Extremely stable over time
 Epigenome	DNA methylation Histone modifications Non-coding RNA	Microarrays Bisulfite sequencing	Continuous data Affected by time and exposures (with reduced plasticity)
 Transcriptome	mRNA	Microarrays RNA sequencing	Continuous data Affected by time and exposures Strong measurement noise
 Proteome	Proteins	Microarrays Mass spectrometry	Continuous data Affected by time and exposures
 Metabolome	Small molecules	Mass spectrometry NMR spectroscopy	Continuous data Structured correlation Strongly affected by exposures
 Microbiome	Microbial DNA	Sequencing	Categorical/Count Data Structured correlation Affected by time and exposures

Motivation

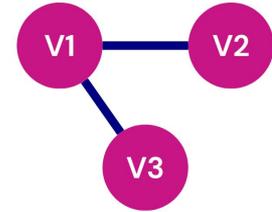
Assumptions

- OMIC variables that are **involved in the same regulatory pathways** tend to be **correlated**
- **Gaussian graphical models** infer **direct associations** between variables by estimating **partial correlations**
[Lauritzen, 1996]

Marginal correlation



Partial correlation



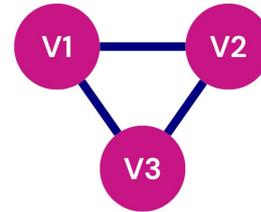
Conditionally on V1, V2 and V3 are independent

Motivation

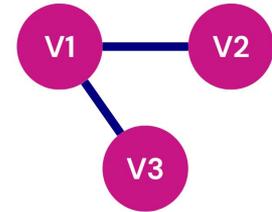
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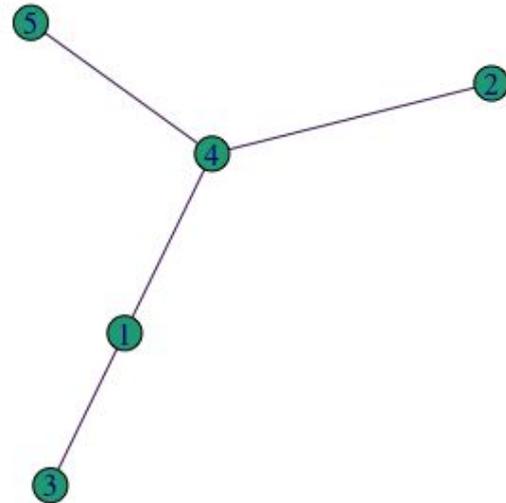
→ These **partial correlations** are encoded in **the precision matrix, Ω**

$$X_i \sim \mathcal{N}(\mu, \Omega^{-1})$$

Motivation

$$\Omega = \begin{pmatrix} 1.71 & 0 & 0.23 & 0.47 & 0 \\ 0 & 1.54 & 0 & -0.54 & 0 \\ 0.23 & 0 & 1.23 & 0 & 0 \\ 0.47 & -0.54 & 0 & 2.52 & 0.5 \\ 0 & 0 & 0 & 0.5 & 1.5 \end{pmatrix}$$

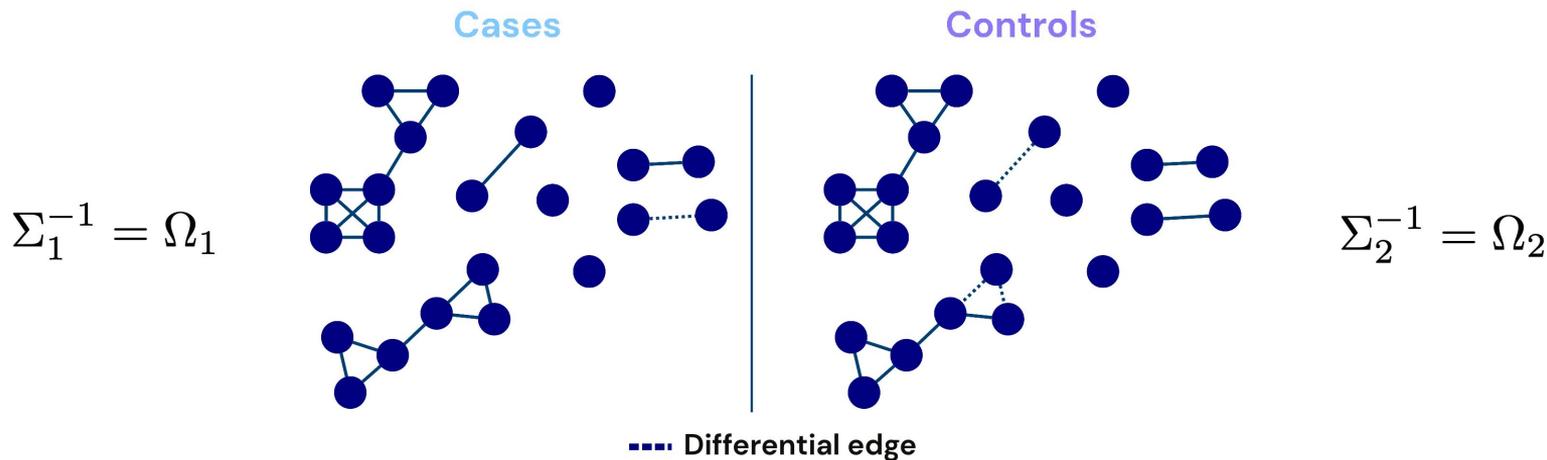
$$A = \begin{pmatrix} 0 & 0 & 1 & 1 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 & 1 \\ 0 & 0 & 0 & 1 & 0 \end{pmatrix}$$



Hypothesis

Let $X \in \mathbb{R}^{n \times p}$ the matrix of observations and $Y \in \{1, 2, \dots, K\}$ the response variable.

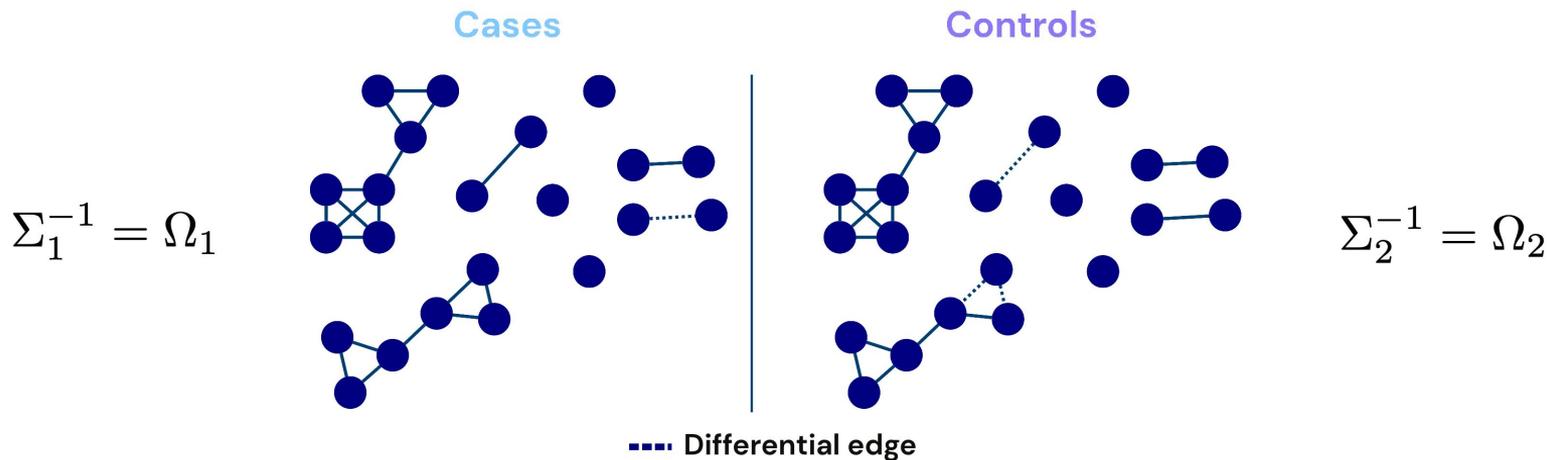
Let $K = 2$, we assume that $X_1 \sim \mathcal{N}(\mu_1, \Sigma_1)$, $X_2 \sim \mathcal{N}(\mu_2, \Sigma_2)$, where $X_1 = (\mathbf{x}_i : Y_i = 1)$ and $X_2 = (\mathbf{x}_i : Y_i = 2)$ such that:



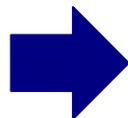
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Graphs have a **shared structure** between classes with **differential edges**



Joint network inference

Research questions

→ Considering joint and independent network estimation approaches, **under what conditions** do they perform best at **recovering shared and condition-specific network structure**?

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Current limitations

- **Various methods** to infer gaussian graphical models and many joint penalties
- Existing studies typically restricted to a **few scenarios**
- Do not clearly distinguish between **model performance** (how well the network structure is recovered) and **model selection** (how penalties are chosen)

Research questions

→ Which **hyperparameter calibration** strategy yields the most reliable edge detection?

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- Do not clearly distinguish between **model performance** (how well the network structure is recovered) and **model selection** (how penalties are chosen)
- Rely on **likelihood-based criteria** (AIC/BIC), which favor dense graphs and reduce interpretability

Joint network inference methods

Method	Loss	Penalties
Graphical lasso (GL) [Yang et al, 2007, Danaher et al, 2014]	$\min_{\{\Omega\}_{k=1}^K} \sum_{k=1}^K n_k \left[\text{tr} \left(\hat{\Sigma}^{(k)} \Omega^{(k)} \right) - \log \det(\Omega^{(k)}) \right]$	<ul style="list-style-type: none"> • L1 penalty on edges • L1 penalty on edges + joint fused penalty • L1 penalty on edges + joint group penalty • L1 penalty on edges + node based penalty
Neighborhood selection (NS) [Meinshausen & Bühlmann, 2006]	$\min_{\{\beta\}_{k=1}^K} \sum_{k=1}^K \sum_{j=1}^p \frac{1}{2n_k} \left\ X_j^{(k)} - X_{-j}^{(k)} B_j^{(k)} \right\ _2^2$ $\beta_{ij}^{(k)} = 0 \Leftrightarrow \omega_{ij}^{(k)} = 0$	<ul style="list-style-type: none"> • L1 penalty on edges • L1 penalty on edges + joint fused penalty • L1 penalty on edges + joint group penalty • Data shared lasso
Partial correlation based neighborhood selection (PC) [Peng et al, 2009]	Let $\tilde{X}_i^{(k)} = \sqrt{\frac{\sigma_{ii}}{\sigma_{jj}}} X_i^{(k)}$ $\min_{\{\rho\}_{k=1}^K} \sum_{k=1}^K \sum_{j=1}^p \frac{1}{2n_k} \left\ X_j^{(k)} - \sum_{i \neq j} \rho_{ij}^{(k)} \tilde{X}_i^{(k)} \right\ _2^2$	<ul style="list-style-type: none"> • L1 penalty on edges • L1 penalty on edges + joint fused penalty • L1 penalty on edges + joint group penalty
Dtrace [Yuan et al, 2017]	Let $\Omega_{\text{diff}} = \Omega_2 - \Omega_1$ $\min_{\Omega_{\text{diff}}} \left[\frac{1}{2} \text{Tr}(\Omega_{\text{diff}}^T \hat{\Sigma}_1 \Omega_{\text{diff}} \hat{\Sigma}_2) - \text{Tr}(\Omega_{\text{diff}}(\hat{\Sigma}_1 - \hat{\Sigma}_2)) \right]$	<ul style="list-style-type: none"> • L1 penalty on differential edges

Penalties

Individual edge penalty

$$\lambda_1 \sum_{k=1}^K \sum_{i \neq j} \left| \Omega_{ij}^{(k)} \right|$$

+

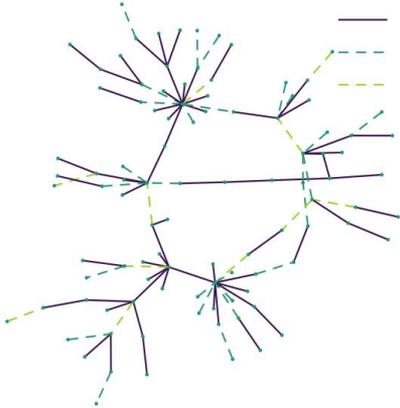
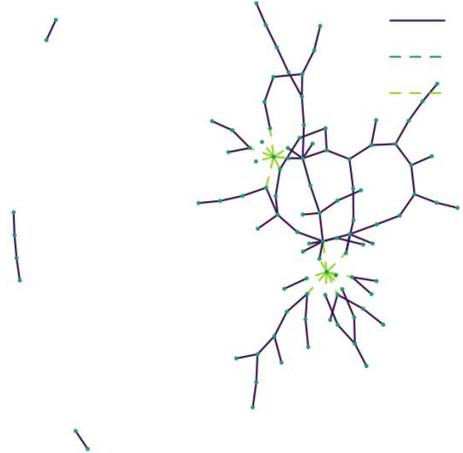
Joint fused penalty

$$\lambda_2 \sum_{k < k'} \sum_{i, j} \left| \Omega_{ij}^{(k)} - \Omega_{ij}^{(k')} \right|$$

Joint group penalty

$$\lambda_2 \sum_{i \neq j} \left(\sum_{k=1}^K \Omega_{ij}^{(k)2} \right)^{1/2}$$

Network generation

Number of nodes, p	30, 100, 300, 500	
Topology	<p style="text-align: center;">Scale-free</p>  <p>— Edge in both (N=64) - - - Only support 1 (N=27) - - - Only support 2 (N=11)</p> <p>The diagram shows a scale-free network with a central hub node connected to many other nodes. The edges are color-coded: solid purple for edges in both, dashed green for edges only in support 1, and dashed yellow for edges only in support 2.</p>	<p style="text-align: center;">Random graph with hubs</p>  <p>— Edge in both (N=88) - - - Only support 1 (N=0) - - - Only support 2 (N=19)</p> <p>The diagram shows a random graph with hubs, where a few nodes have a high degree of connectivity. The edges are color-coded: solid purple for edges in both, dashed green for edges only in support 1, and dashed yellow for edges only in support 2.</p>
Differential edge	<p>→ Random (a percentage of edges turned off)</p> <p>→ Hub-based</p>	

Data generation

- Generate **2 positive definite matrices** for the corresponding graphs
- Σ_1, Σ_2 : computed as the **inverse of precision matrices**
- We simulate data from a **multivariate Normal distribution** $N(0, \Sigma_k)$, where $k \in \{1, 2\}$
- Sample sizes considered: **N = 50, 100** per condition
- **50 independent replicates** of datasets X_1 and X_2

Performance evaluation

Evaluation of differential support edges

Compute over λ_2 :

- Area under the precision-recall curve (AUPR)
- Area under recall vs. false-positive rate (AUC)
- Area under power vs. FDR (AUPF)

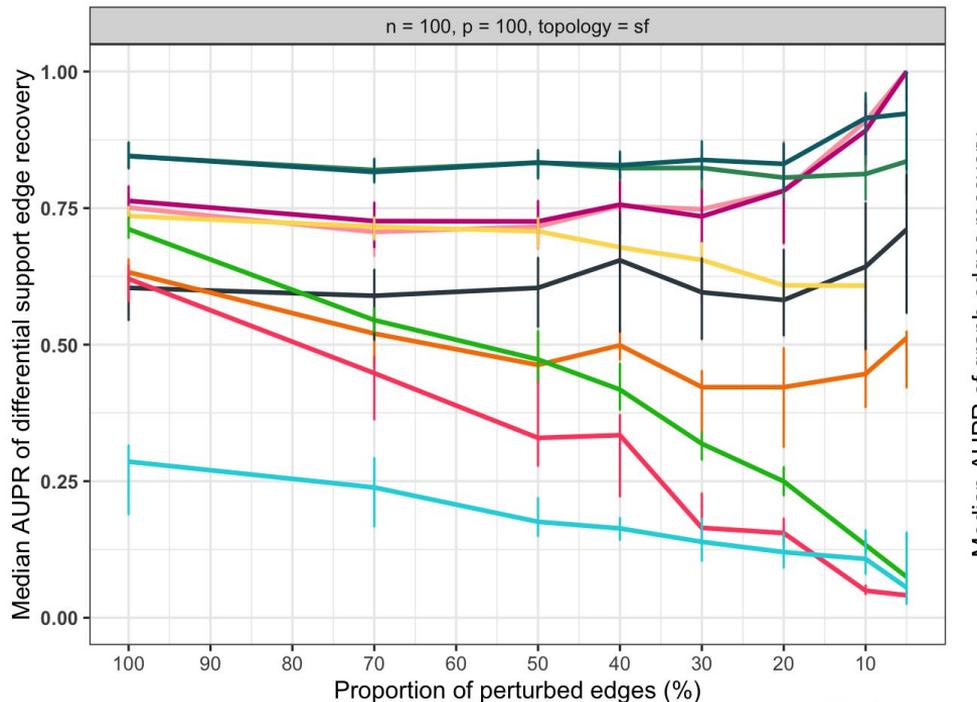
Evaluation of overall graph inference

- Apply the same metrics to assess graph recovery over λ_1

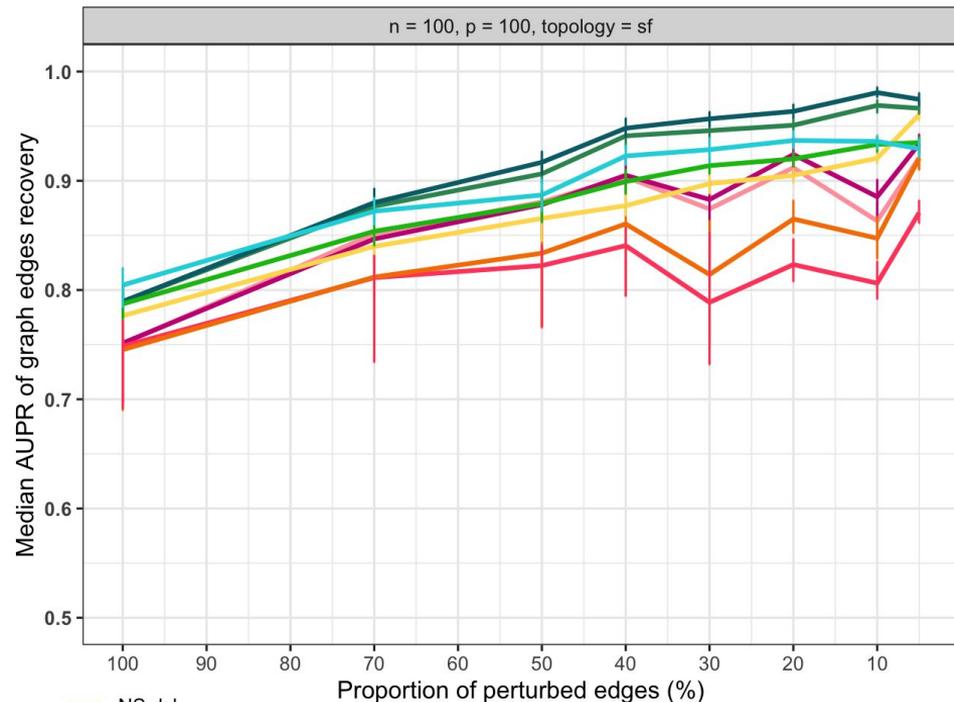
For each method, record the **maximum F1-score** in terms of edge recovery across λ_1 , λ_2 parameters and average it over the 50 simulated datasets

Results: random edge perturbations

Differential support recovery

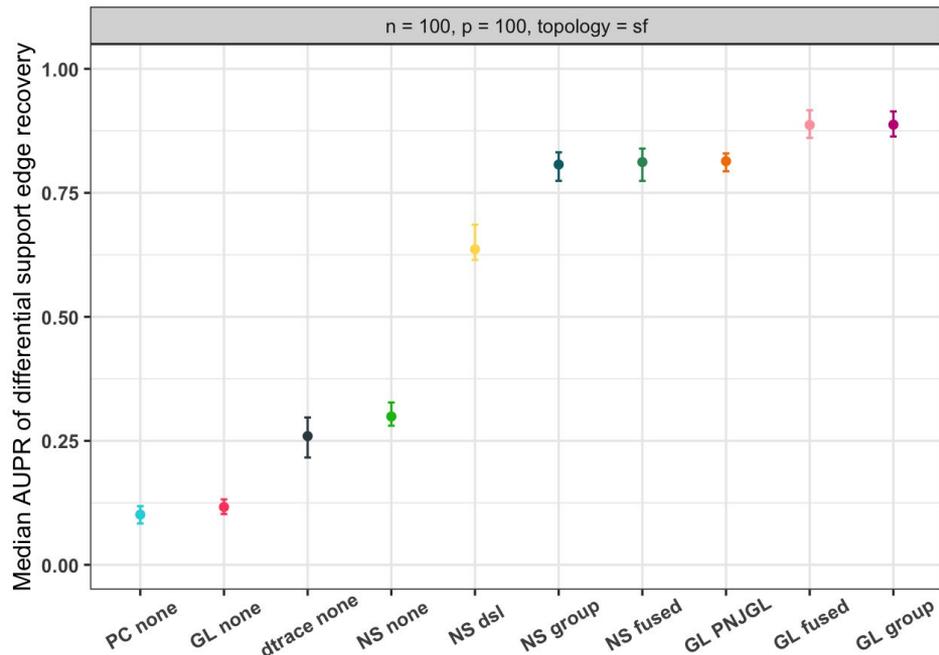


Graphs support recovery

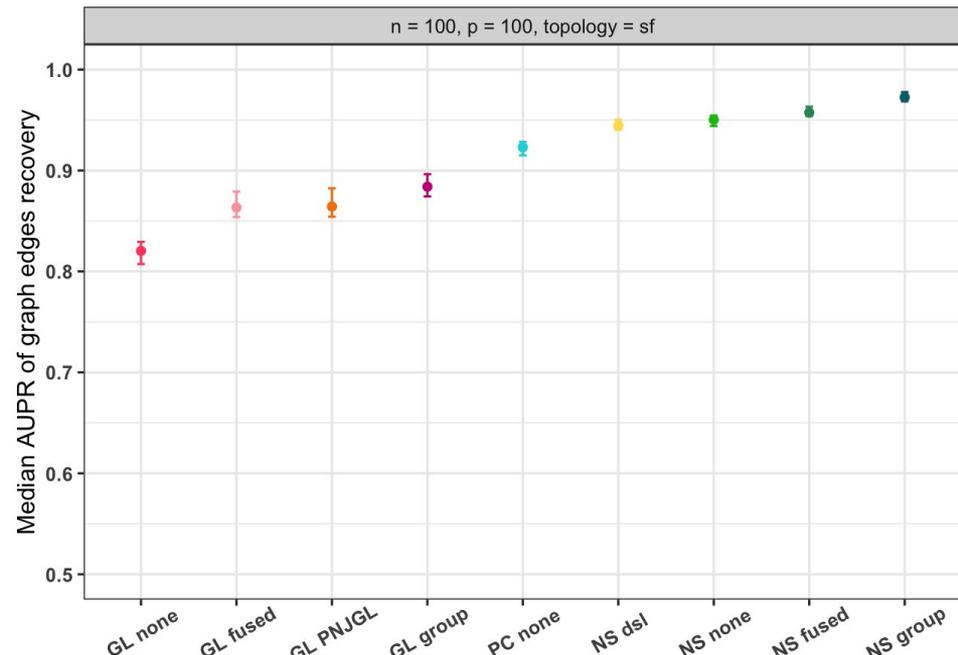


Results: hub based perturbations

Differential support recovery



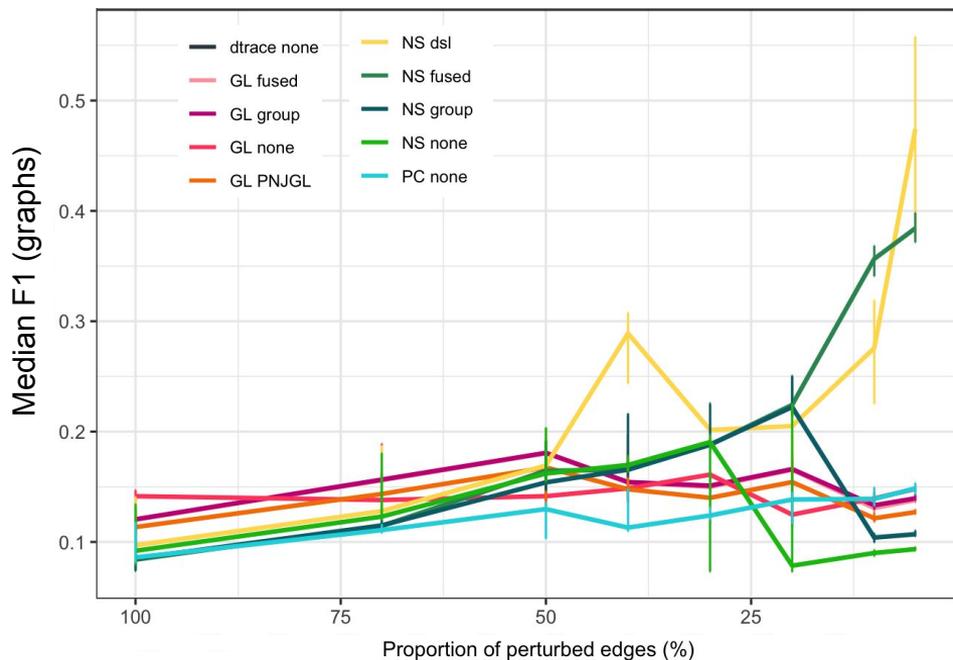
Graphs support recovery



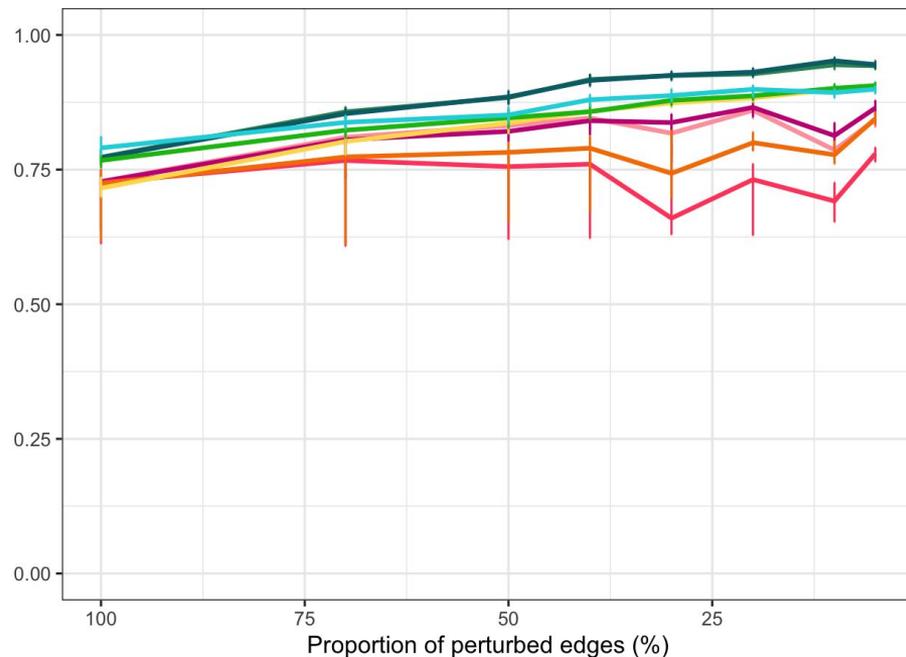
Model calibration

Often relies on **likelihood-based criteria (AIC/BIC)**, which favor **dense graphs** and reduce interpretability

Calibration with AIC



Oracle



Model calibration

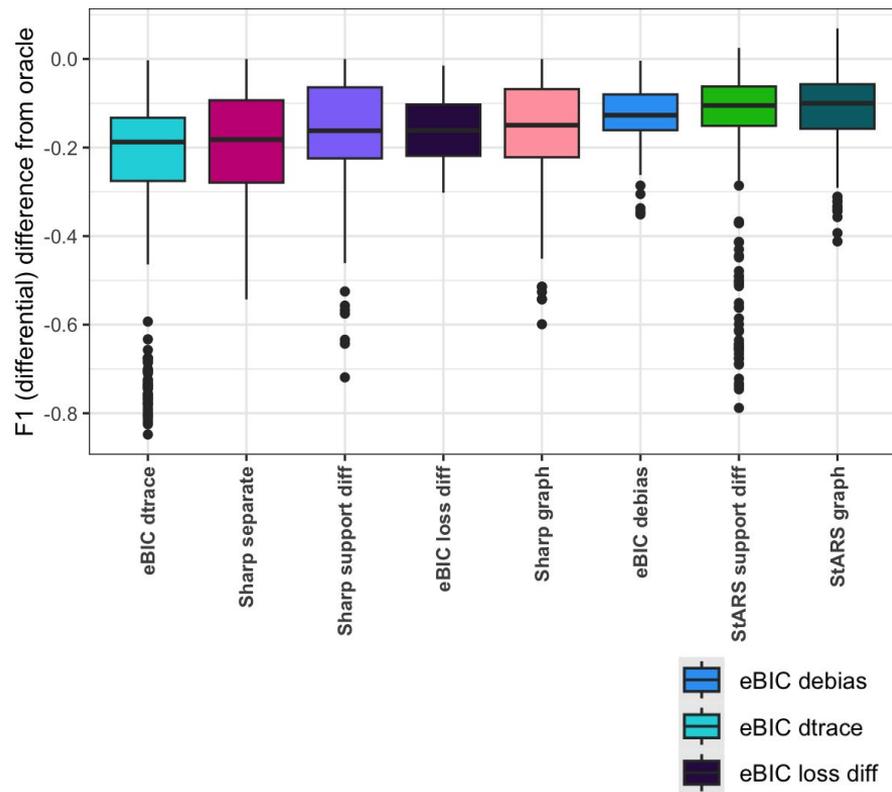
1. Obtain debiased precision matrices

2. Calibrate

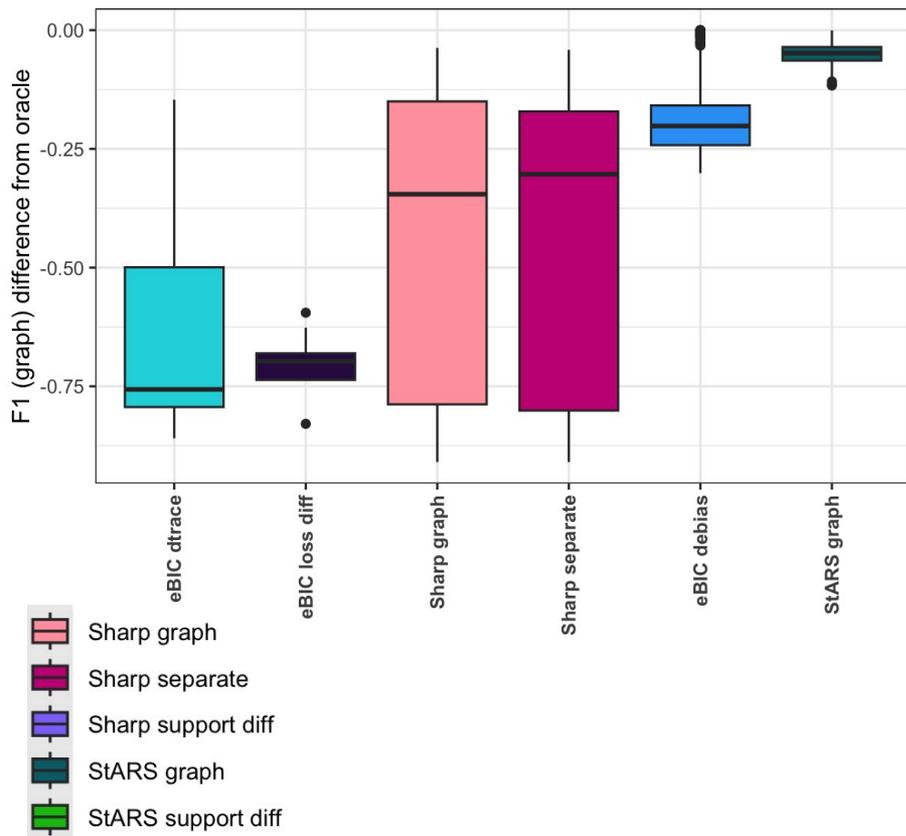
- $\text{eBIC}(\lambda_1, \lambda_2) = \sum_{k=1}^K n_k \left(-\log \det(\hat{\Omega}^{(k)}) + \text{tr}(\hat{\Sigma}^{(k)} \hat{\Omega}^{(k)}) \right) + \log(n_k) \frac{\|\hat{\Omega}^{(k)}\|_0}{2} + 4\gamma \log(p) \frac{\|\hat{\Omega}^{(k)}\|_0}{2}$
- **Dtrace pseudo eBIC** $(\lambda_1) = \left[\frac{1}{2} \text{Tr}(\hat{\Omega}_{\text{diff}}^T \hat{\Sigma}_1 \hat{\Omega}_{\text{diff}} \hat{\Sigma}_2) - \text{Tr}(\hat{\Omega}_{\text{diff}}(\hat{\Sigma}_1 - \hat{\Sigma}_2)) \right]$
 $+ \log(n_1 + n_2) \|\hat{\Omega}_{\text{diff}}\|_0 + 4\gamma \log(p) \|\hat{\Omega}_{\text{diff}}\|_0$
- **Stability selection** (StARS [Liu et al, 2010], Sharp [Bodinier et al, 2023])

Model calibration

Differential support recovery



Graphs support recovery



Weights

→ Improving edge selection and incorporating prior knowledge

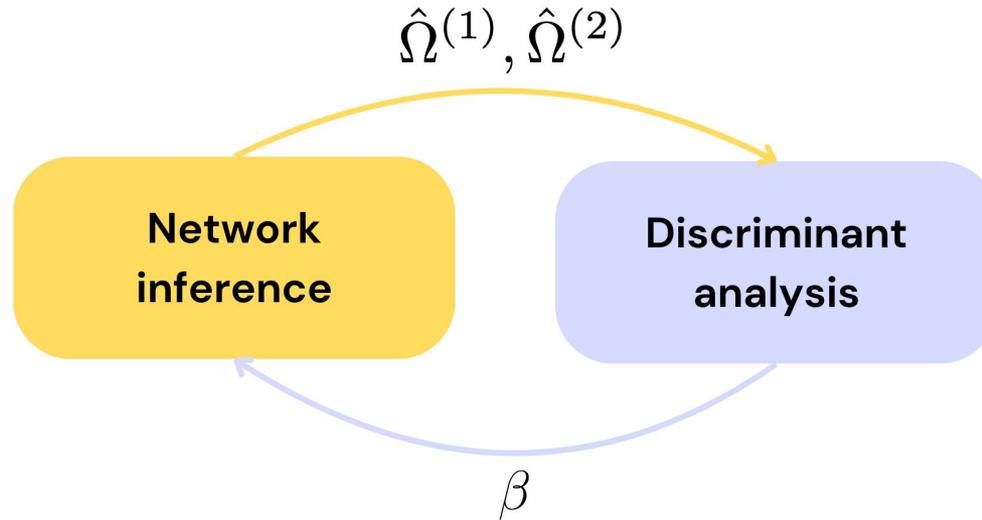
	Individual edges	Joint fused	Joint group
Penalty	$\lambda_1 \sum_{k=1}^K \sum_{i,j} w_{ij}^{(k)} \left \beta_{ij}^{(k)} \right $	$\lambda_2 \sum_{k < k'} \sum_{i,j} w_{ij}^{(kk')} \left \beta_{ij}^{(k)} - \beta_{ij}^{(k')} \right $	$\lambda_2 \sum_{i,j} w_{ij} \left(\sum_{k=1}^K \beta_{ij}^{(k)2} \right)^{1/2}$
Adaptive weight	$w_{ij}^{(k)} = \frac{1}{\left \beta_{ij}^{(k)} \right }$	$w_{ij}^{(kk')} = \frac{1}{\left \beta_{ij}^{(k)} - \beta_{ij}^{(k')} \right }$	$w_{ij} = \frac{1}{\left(\sum_{k=1}^K \beta_{ij}^{(k)2} \right)^{1/2}}$
Degree-based weight	$w_{ij}^{(k)} = \frac{1}{\sqrt{(d_i^{(k)} + 1) \cdot (d_j^{(k)} + 1)}}$	$w_{ij}^{(kk')} = \frac{1}{\left \sqrt{(d_i^{(k)} + 1) \cdot (d_j^{(k)} + 1)} - \sqrt{(d_i^{(k')} + 1) \cdot (d_j^{(k')} + 1)} \right + \epsilon}$	$w_{ij} = \frac{w_{ij}^{(k)} + w_{ij}^{(k')}}{2}$

Weights

- **Improving edge selection** and incorporating **prior knowledge**
 - ◆ Weights obtained from **Stochastic Block Model** clustering on the adjacency matrix [Ambroise et al, 2009]
 - ◆ Weights obtained from **Multipartite Block Model** to account for multi-omics [unpublished Chiquet, 2018]

Next steps

- Combine joint network estimation and outcome discrimination to:
 - ◆ improve **interpretability** and **discrimination**
 - ◆ Identify both **structural** and **predictive differences** across classes





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Thank you

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